Inversions in split trees and conditional Galton–Watson trees

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- 1. The definition
- 2. Inversions in fixed trees
- 3. Inversions in split trees
- 4. Inversions in conditional Galton-Watson trees

The definition

- Let σ_1,\ldots,σ_n be a permutation of $\{1,\ldots,n\}.$
- If i < j and $\sigma_i > \sigma_j,$ then the pair (σ_i, σ_j) is called an inversion.



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- Let T be a tree with node set V.
- Let λ be node labeling $\lambda:V\to\{1,\ldots,|V|\}.$
- Define the number of inversions

$$I(T,\lambda) \stackrel{\mathrm{def}}{=} \sum_{u < v} \mathbf{1}_{\lambda(u) > \lambda(v)}.$$



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- Flajolet, Poblete, and Viola (1998) showed that this random variable for Cayley trees converges to an Airy distribution.
- Panholzer and Seitz (2012) generalized this to conditional Galton–Watson trees.

Note that

$$\mathbb{E}\left[I(T)\right] = \sum_{u < v} \mathbb{E}\left[\mathbf{1}_{\lambda(u) > \lambda(v)}\right] = \frac{1}{2} \sum_{u < v} 1 \stackrel{\text{def}}{=} \frac{1}{2} \Upsilon(T).$$

• $\Upsilon(T)$ is also known as the *total path length*, since

$$\Upsilon(T) \stackrel{\text{def}}{=} \sum_{v} d(v),$$

where d(v) is the depth of v.

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Simulations

- We did simulation on the complete binary tree of height 26
- Does the result suggest a central limit theorem?



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So it cannot be a normal distribution!

A key observation

- Let Z_o be the number of inversions involving the root.
- Then Z_o and the numbers of inversions in the left subtree and right subtree are independent.
- Proof by conditioning on the labels that go the left and the right.



- Let \boldsymbol{z}_v be the size of the subtree at $\boldsymbol{v}.$
- Let ${\cal Z}_v$ be the number of inversions involving v and one of its descendants.

Lemma 1

Let T be a fixed tree. Then

$$I(T) \stackrel{\mathrm{d}}{=} \sum_{v \in V} Z_v,$$

where $\{Z_v\}_{v \in V}$ are independent random variables, and $Z_v \sim \text{Unif}\{0, 1, \dots, z_v - 1\}.$

Cumulants

• The cumulant-generating function of a r.v. X is

$$K_X(t) = \log \mathbb{E}\left[e^{t\,X}\right].$$

- The cumulants $\varkappa_k(X)$ are defined by

$$K_X(t) = \sum_{k \geq 1} \varkappa_k(X) \frac{t^k}{k!}.$$

• If X is independent of Y, then

$$\varkappa_k(X+Y)=\varkappa_k(X)+\varkappa_k(Y).$$

• We can compute centralized-moments from cumulants.
k-total common ancestors

- For k nodes $v_1,\ldots,v_k,$ let $c(v_1,\ldots,v_k)$ be the number of ancestors that they share.
- We define

$$\Upsilon_k(T) \stackrel{\mathrm{def}}{=} \sum_{v_1, \dots, v_k} c(v_1, \dots, v_k).$$

 $\bullet \quad \text{Note that } \Upsilon(T) = \Upsilon_1(T) - |V|.$



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- Note that $\Upsilon(T)=\Upsilon_1(T)-|V|.$



Theorem 2

We have

$$\begin{split} \mathbb{E}\left[I(T)\right] &= \varkappa_1(I(T)) = \frac{1}{2}(\Upsilon_1(T) - |V|),\\ \mathrm{Var}\left(I(T)\right) &= \varkappa_2(I(T)) = \frac{1}{12}(\Upsilon_2(T) - |V|). \end{split}$$

More generally, for $k \ge 1$,

$$\varkappa_{2k+1}(I(T)) = 0, \qquad \varkappa_{2k}(I(T)) = \frac{B_{2k}}{2k}(\Upsilon_{2k}(T) - |V|),$$

where \boldsymbol{B}_k denotes the k-th Bernoulli number.

Theorem 3

Let ${\cal T}_n$ be a sequence of fixed trees on n nodes. Let

$$X_n = \frac{I(T_n) - \mathbb{E}\left[I(T_n)\right]}{\sqrt{\Upsilon_2(T_n)}}.$$

Assume that for all $k \geq 1$,

$$\frac{\Upsilon_{2k}(T_n)}{\Upsilon_2(T_n)^k} \to \zeta_{2k},$$

for some sequence (ζ_{2k}) . Then there exists a unique X with

$$\varkappa_{2k-1}(X)=0,\qquad \varkappa_{2k}(X)=\frac{B_{2k}}{2k}\zeta_{2k},\qquad k\geq 1,$$

such that $X_n \stackrel{\mathrm{d}}{\longrightarrow} X$.

• Let $T_n = P_n$ (a path of length n),

$$\Upsilon_k(T_n)\sim \frac{1}{k+1}n^{k+1}.$$

Thus

$$\frac{\Upsilon_{2k}(T_n)}{\Upsilon_2(T_n)^k} \to 0 \qquad (k\geq 2).$$

- So $(I(T_n) - \mathbb{E}\left[I(T_n)\right])/n$ converges to X with

$$\varkappa_k(X)=0, \qquad (k\geq 3).$$

• Then X must be a normal distribution.

Theorem 4

Let $b\geq 2$ and let T_n be the complete $b\mbox{-}{\rm ary}$ tree of height m with n nodes. Then

$$X_n = \frac{I(T_n) - \mathbb{E}\left[I(T_n)\right]}{n} \overset{\mathrm{d}}{\longrightarrow} \sum_{d \geq 0} \sum_{j=1}^{b^d} \frac{U_{d,j}}{b^d},$$

where $(U_{d,j})_{d \ge 0, j \ge 1}$ are independent Unif[-1/2, 1/2].

Inversions in split trees

- BST is a computer data structure for storing "item" according to the order of their "keys".
- BST can be defined with a bijection to permutations.
- The average height of a BST of size n is $\alpha \ln n \beta ln lnn$ [Reed, 2003].



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- We can construct BST in another way.
- Consider an infinite binary tree.
- Each node is a "bucket" of size one.
- Each node is given a split vector $\mathcal{V} = (U, 1-U)$ chosen independently.
- *n* balls come into the root one by one.
- When a bucket has more than one node, the extra goes to child nodes chosen at random according to V.
- All empty buckets are removed in the end.



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Split trees

- By choosing different:
 - inifnite trees
 - bucket sizes
 - distributions of split vector
- split trees encompasses:
 - binary search trees
 - b-ary search trees
 - digital search trees
 - tries, etc.
- Split trees are introduced by Devroye (1999).

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- We first choose the split tree ${\cal T}_n$ with n balls.
- Then we randomly label the balls.
- We define $\widehat{I}(T_n)$ as the number of inversions for balls.
- We study

$$\hat{X}_n = \frac{\hat{I}(T_n) - \mathbb{E}\left[\hat{I}(T_n)\right]}{n}$$

Theorem 5

Т

Let T_n be a b-ary split tree with bucket size $s_0.$ Let $\mathcal{V}=(V_1,\ldots,V_b)$ be a split vector. Let \hat{X} be the unique solution for the fixed-point equation

$$\hat{X} \stackrel{\rm d}{=} \sum_{i=1}^{b} V_i \hat{X}^{(i)} + \sum_{j=1}^{s_0} U_j + \frac{s_0}{2} D(\mathcal{V}).$$
 hen $\hat{X}_n \stackrel{\rm d}{\longrightarrow} \hat{X}.$

- Proof by the contraction method.
- A similar result holds for labeling nodes instead of balls.

Inversions in conditional Galton-Watson trees

- A Galton–Watson tree starts with a root node.
- Each node in the tree is given a random number of child nodes.
- The numbers of children are independent with distribution ξ .

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- A conditional Galton–Watson tree ${\cal T}_n$ is a Galton–Watson tree conditioned on having n nodes.
- It encompasses (uniform random)
 - plane trees
 - binary trees
 - b-ary trees
 - Cayley trees
- Very well studied, see, e.g., Janson (2012).

The depth-first walk on a conditional Galton–Watson tree



It is well-known that the depth-first walk on conditional GW trees converges to Brownian excursions [Aldous (1991a), Aldous (1991b), Aldous (1993), and Le Gall (2005)].

Let e(t) be a Brownian excursion. Let

$$\eta \stackrel{\text{def}}{=} 4 \int_{0 \le s \le t \le 1} \min_{s \le u \le t} e(u) \mathrm{d}s \mathrm{d}t.$$

Theorem 6

Assume that $\mathbb{E}[\xi] = 1$, $\operatorname{Var}(\xi) = \sigma^2 \in (0, \infty)$, and $\mathbb{E}[e^{\alpha \xi}] < \infty$ for some $\alpha > 0$. Then

$$\frac{I(T_n) - \frac{1}{2} \Upsilon(T_n)}{n^{5/4}} \xrightarrow{\mathrm{d}} \frac{1}{\sqrt{12\sigma}} \sqrt{\eta} \ N(0, 1).$$

Results for conditional GW trees ii

Let

$$X_n = \frac{I(T_n) - \mathbb{E}\left[I(T_n)\right]}{n^{3/2}}$$

Then we can decompose

$$X_n = \frac{I(T_n) - \frac{1}{2}\Upsilon(T_n)}{n^{3/2}} + \frac{\Upsilon(T_n) - \mathbb{E}\left[\Upsilon(T_n)\right]}{2n^{3/2}}$$

- Our result shows that the first term goes to zero.
- Aldous (1991b) showed that $\Upsilon(T_n)/n^{-3/2}$ converges to an Airy distribution.
- So X_n also converges to an Airy distribution.
- We recover result from Panholzer and Seitz, 2012.

- Let σ be a permutation of $\{1,\ldots,k\}.$ Let

$$R_{\sigma}(T,\lambda) = \sum_{u_1 < \cdots < u_k} \mathbf{1}_{[\lambda(u_1, \dots, u_k) = \sigma]}.$$

• Then
$$R_{21}(T,\lambda) = I(T,\lambda).$$

- Recently Albert, Holmgren, Johansson, and Skerman, 2018 studied $R_{\sigma}(T,\lambda)$ for complete binary trees and split trees.
- The spirit most questions about permutations can be asked for trees.

Questions?



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